- A comparison of statistical and deterministic methods for shallow landslide susceptibility zoning in
 clayey soils
- 3 Authors: Ciurleo M., Cascini L., Calvello M.
- 4

5 ABSTRACT

Shallow landslides are widespread in different geological contexts and generally occur as multiple 6 7 events over large areas. When these phenomena involve fine-grained soils, they may cause serious 8 consequences—in terms of environmental and property damages—and thus their spatial forecasting 9 becomes a relevant issue for land use planning and design purposes. The existing literature provides 10 several methods for landslide susceptibility assessment, categorized in qualitative and quantitative methods. When dealing with analyses at large scale (1:5000), quantitative methods are generally 11 preferred. In this paper, landslide susceptibility maps are produced for a study area prone to shallow 12 landsliding, located in Catanzaro (southern Italy). To this aim, two quantitative methods are 13 implemented: the statistical "information value method" and the deterministic "TRIGRS model." 14 15 The two approaches are compared by means of two indicators of the grade of correctness of the landslide susceptibility maps: the area under curve of the ROC curve, AUC, and the overestimation 16 17 index, OI. The results of the analyses in terms of AUC values demonstrate the effectiveness and consistency of both methods in performing the susceptibility mapping of the study area. When the 18 19 OI values are considered, the results provided by the deterministic model are slightly better than the ones resulting from the statistical analysis. This does not come as a surprise for the case study at 20 21 hand and it can be ascribed to the availability, within the study area, of: a reliable database of soil properties, and an in-depth knowledge of the behaviour of the considered landslides. 22

23

Keywords: fine-grained soils, shallow landslides, susceptibility, zoning, statistical analysis,
deterministic analysis, comparison

26

27 Highlights

- Susceptibility zoning for shallow landslides in fine-grained soils at large scale
- Statistical and deterministic methods applied to a study area in Italy
- Comparison of zoning maps from statistical and deterministic analyses
- Proven effectiveness of both analyses for susceptibility zoning at large scale
- 32

33 1. Introduction

Rainfall-induced shallow landslides are widespread all over the world and, during a single rainfall event, they can involve large areas (Park et al., 2013; Lee and Park, 2016; Romer and Ferentinou, 2016; among others). The consequences caused by these phenomena are linked to a series of factors, among which: the geological context of the area affected by landsliding, the mechanical characteristics of the soils, the vulnerability of the exposed elements.

Shallow landslides of flow type in coarse-grained soils are generally characterized by both scarcity 39 of warning signs in the pre-failure stage and high velocities in the post-failure phase (McDougall 40 41 and Hungr, 2004; Sorbino et al., 2010; Hungr et al. 2014; Yerro et al., 2016). On the other hand, shallow landslides involving clayey colluvial soils frequently present warning signs in the pre-42 failure stage (e.g., tension cracks at the top of the slope) and a shorter run-out in the post failure 43 44 phase (Hungr et al. 2001; Meisina et al., 2006; Cascini et al., 2015). Many examples of rainfall-45 induced shallow landslides in both coarse-grained and fine-grained soils are reported in Europe (e.g., Borrelli et al., 2012; Martinović et al., 2016), America (e.g., Baum et al., 2005; Godt et al., 46 47 2008), Asia (e.g., Park et al., 2013; Hadmoko et al., 2017) and Africa (e.g., Broothaerts et al., 2012; Romer and Ferentinou, 2016). A significant number of contributions specifically deal with the 48 assessment of shallow landslide susceptibility (e.g., Leventhal and Kotze, 2008; Nandi and 49 Shakoor, 2009; Frattini et al., 2010; Kavzoglu et al., 2015; Romer and Ferentinou, 2016). 50

Landslide susceptibility assessment, herein intended as the landslide spatial probability of 51 occurrence within a given territory, can be performed by qualitative or quantitative methods, 52 depending on how the landslide causal factors are considered and modelled (Lee and Park, 2016). 53 54 The applicability of each method depends on: the availability, quality and accuracy of the data; the resolution of zoning; the required outcomes; and the scale of analysis (Soeters and van Westen, 55 1996; Cascini, 2008; Fell et al., 2008). Quantitative analyses can be divided into statistical and 56 deterministic methods (Soeters and van Westen, 1996). Statistical methods, applicable at different 57 scales (from 1:100,000 to 1:5000), establish the relationships between predisposing factors and 58 59 landslides through the proper use of statistic indicators (Carrara, 1983; Baeza and Corominas, 2001; Nefeslioglu et al. 2008), while neglecting to explicitly model the landslide failure mechanisms 60 (Park et al., 2013). Deterministic methods, applicable at large and detailed scales ($\geq 1:5000$), 61 properly analyze existing or potential failure mechanisms via physically-based models calibrated 62 using on-site and laboratory test results (Salciarini et al., 2006; Huang and Kao 2006; Godt et al., 63 64 2008; Park et al., 2013).

This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 2)

Several significant examples of the application of landslide statistical analyses at large scale are 65 available in the literature (Cervi et al., 2010; Reza and Daneshvar, 2014; Iovine et al., 2014; Regmi 66 et al., 2014). In some cases, a separation between different types of landslides is lacking and, for 67 instance, slides, creep phenomena and falls may be considered together when deriving the event 68 map for the statistical correlations (Hadmoko et al., 2017). On the other hand, a proper 69 understanding of the landslide triggering processes is typically a key prerequisite for a consistent 70 71 application of physically based models estimating shallow rainfall-triggered landslide susceptibility (Godt et al., 2008; Sorbino et al., 2010; Cascini et al., 2017). Only few contributions deal with the 72 73 comparison between the results of statistical and deterministic methods applied to the same study 74 area at large scale (Cervi et al. 2010; Armas et al., 2013).

75 Starting from the abovementioned aspects, the paper highlights the important role played by the understanding of the landslide mechanisms for the susceptibility assessment of shallow landslides 76 77 in clayey soils. This aim is pursued by means of a skilled application of both statistical and deterministic methods in a study area located in southern Italy. The issue is dealt with at large scale 78 79 (1:5000) thus allowing a thorough study of all the factors playing a significant role during the prefailure and failure stages of landsliding. Particularly, the paper preliminary focuses on the 80 81 characteristics of the analysed phenomena, then performs a susceptibility assessment of the study area by means of the statistical method known as "information value" (e.g. Yin and Yan, 1988) and 82 of the physically-based model "TRIGRS" (e.g. Montgomery and Dietrich, 1994). Both methods end 83 with the production of landslide susceptibility computational maps. The maps are compared and 84 critically analysed in order to highlight advantages and limitations related to the two different 85 methods of analysis. 86

87

88 2. Materials and Method

89 2.1. Study area and data set

The study area is located in southern Italy, in the Calabria region, within a larger area of 136 Km² (Fig. 1) that has been frequently affected by landslides (Antronico et al., 2013; Borrelli et al., 2015; Gullà et al., 2004, 2008; Cascini et al., 2015; Ciurleo et al., 2016). In this area, more than 60% of not flat land (i.e. slope angle greater than 5°) is covered by Pliocene light blue-grey silty clays. In this area two sites, indicated in the Figure as S2 and S3, have been already analysed by Gullà et al. (2004), Cascini et al. (2015) and Cascini et al. (2017). The focus of the present paper is the site, indicated in the Figure as S1, already partly analysed by Cascini et al. (2015) and Calvello and 97 Ciurleo (2016). The chosen site of interest is deemed to be the most representative in relation to two
98 important shallow landslide events, which respectively occurred in 2009 and 2010.

The homogeneity of the fine-grained soils outcropping in the sites S1, S2 and S3 emerges from a 99 comparison of geotechnical properties (Fig. 2). In the entire area, the soils can be classified as 100 inorganic, inactive clays characterized by high plasticity and a high liquid limit (Fig. 2b). All the 101 grain size distribution data fall within a well-defined grain size envelope; the upper limit shows a 102 fine-grained fraction ranging from 89% to 99%, while the amount of sand varies from 1% to 11%. 103 The index properties of the soil, the minimum and maximum values assumed by the void ratio and 104 105 the soil porosity are summarized in Fig. 2b. The available shear strength envelope ranges from an 106 upper limit, with a cohesion value of 24.3 kPa and a friction angle of 35.2°, to a lower limit 107 characterized by a cohesion value of 2 kPa and a friction angle of 22.3°. The saturated permeability (Ks) ranges from 3.1×10^{-8} m/s to 7.65×10^{-7} m/s (Gullà et al., 2004, 2008; Cascini et al., 2015), 108 reaching a value of 5×10^{-6} m/s in the topmost cracked weathered layers (Cascini et al., 2017). 109

Owing to the homogeneity of the geotechnical data, the information gathered by previous studies in 110 111 the sites S2 and S3 have been combined with those available in the study areas, i.e. site S1. The study area is located on the left bank of the Corace river, it covers approximately 8 km² and it 112 represents the product of a complex geological evolutive model provided by Cascini et al. (2015). 113 The topographical and morphological data used for the analyses are derived from a digital elevation 114 model with a 5 m resolution generated from digital topographic maps of the Calabria Region 115 (1:5000 scale, year 2005). The available geological information, provided in Fig. 3, shows that the 116 outcropping lithology is mainly characterized by silty clays, partially affected by intercalations of 117 sands. A sandstone layer constituting a morphoselective scarp, which is about 10 to 30 m thick and 118 119 covers 2.5% of the test site, will not be included in the analysis, given the study focuses on shallow landslide susceptibility in fine-grained soils. 120

The fine-grained soils, outcropping in S1 and in nearby areas, are involved in several processes that can be considered landslide predisposing factors (e.g., weathering of the most shallow layers, formation of cracks). The combination of these factors leads to an evolutionary pattern of the landslides lasting few years, from the first warning signs in the pre-failure stage to the end of movements (Fig. 4). It follows that the landslide inventory map has to be continuously updated. As reported in Fig. 5a, the original landslide inventory in the study area is dated 2009 and was updated in 2010.

Using aerial photos and satellite images from Google Earth, 117 and 509 shallow landslide triggering areas of slide type were inventoried in 2009 and in 2010, respectively. It is worth

highlighting that in 2009 the total area affected by landslides is equal to 0.09 Km², while in 2010 130 this area is about 0.2 Km^2 . The smallest triggering landslide area is about 9 m² while the largest one 131 is 6000 m²; the depths of the slip surfaces range from few decimetres to 3 meters. Considering that 132 the sliding surfaces' depths depend on their location on the slope, Cascini et al. (2017) proposed a 133 methodology to map the thickness of degraded rock. The methodology, which was initially 134 calibrated and validated by the authors over one morphological hollow (0.3 km² size), is herein 135 applied to map the weathered thickness for the entire study area (Fig. 5b). Fig. 5b underlines the 136 137 presence of low values of thickness on sharply defined ridges (up to 0.5 m) and at the top of the 138 slopes (from 0.5 m to 1.5 m) and maximum thickness depths higher than 5m at the valley bottom.

Significant phases of the morphological slope evolution due to shallow landslides are generally 139 140 related to short and intense rainfall events. The analysis of the rainfall patterns in relation to the activity of the considered landslides is beyond the scope of the present paper yet it's worth 141 142 highlighting, herein, some findings on "critical rainfall" conditions for the area coming from previous works (Gullà et al., 2004; Cascini et al., 2017). Referring to shallow landslides that 143 144 respectively occurred in site S2, in 2002 and 2003, and in Site S1, in 2009, these studies indicated the maximum intensity of the rainfall event ranging from 43 mm (2002-2003) to 70 mm (2009) over 145 146 2 consecutive days and equal to 162 mm over six consecutive days (2009).

147

148 2.2 Methods

The methodology used herein to assess the susceptibility to shallow landslides in fine-grained soils starts with the calibration of the statistical and physically-based models and it ends with the production of landslide susceptibility computational maps by means of both statistical and deterministic methods. Both maps are computed and plotted using terrain computational units, TCUs (Calvello et al., 2013), and discretizing the study area by means of square cells whose dimension (herein 5 m by 5 m) is related to the employed scale of the analysis (1:5000 scale).

The statistical analyses are based on bivariate correlations (Tangestani, 2009; Conforti et al., 2012; Ciurleo et al., 2016) over the study area, between available independent variables (e.g. elevation zone, slope gradient, slope curvature) and a dichotomous dependent variable derived from the available landslide inventories. The independent variables can be either categorical or numerical. The categorical variables are divided in a number of classes directly correlated to the subdivision in classes of the thematic maps from which they are derived; the numerical variables are herein always divided in eight classes using the the Jenks Natural Breaks algorithm (Jenks, 1977). Following this 162 classification method, class breaks are identified as the boundaries that best group similar values163 and that maximize the differences between classes.

The dependent variables used for the analyses are the two already-mentioned landslide inventories dated 2009 and 2010. For both landslide inventories, the statistical weight assigned to each class, j, of each variable, Vi, is computed using the following formula, originally proposed within the "information value method" (e.g. Yin and Yan, 1988):

168

169
$$W_{ij} = \log\left(\frac{D_{ij}}{D^*}\right) = \log\left(\frac{F_{ij}/N_{ij}}{F_{tot}/N_{tot}}\right)$$
(1)

170

where: W_{ij} is the weight assigned to the class j of the independent variable Vi; D_{ij} is the density of landslides within class j of the independent variable Vi; D^* is the average density of landslides within the study area; F_{ij} is the number of TCUs with landslides belonging to the class j of the independent variable Vi; N_{ij} is the number of TCUs belonging to the class j of the independent variable Vi; F_{tot} is the total number of TCUs with landslides within the study area; N_{tot} is the total number of TCUs with landslides within the study area; N_{tot} is the total number of TCUs within the study area.

When W_{ij} assumes low negative value, the statistical implication is low probability for TCUs belonging to a given class of a given variable to be affected by landslides; when W_{ij} assumes high positive values, there's a high probability that TCUs belonging to that class are affected by landslides. When landslides are not present in a given class of a given independent variable, Eq. 1 cannot be used to compute the weight values. In such cases, the class weight is herein set to a value equal to the closest negative integer inferior to the minimum computed weight for all classes of all variables (Ciurleo et al., 2016).

184 The performance assessment of the bivariate correlation between the independent and the dependent 185 variables used herein is based on the criteria proposed by Ciurleo et al. (2016), which employ two 186 indicators, β and σ , defined as follows:

187

188
$$\beta_i = \frac{TPR_i}{FPR_i} = \frac{Sensitivity_i}{1-Specificity_i}$$
(2)

190
$$\sigma_i = \sqrt{\frac{\sum_{j=1}^n (W_{ij}^* - \underline{W}_i)^2}{n-1}}$$
 (3)

191

This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 6)

where: TPR_i is the true positive rate (sensitivity of the bivariate model) for the independent variable Vi; FPR_i is the false positive rate for the independent variable Vi (1- specificity of the bivariate model); W_{ij}^* is the normalized value of the weight assigned to the class j of the independent variable Vi (Ciurleo et al., 2016); W_i is the average value of the weights assigned to the classes of the independent variable Vi; n is the number of classes of the independent variable Vi.

197 The two indicators computed with Eqs. (2) and (3) have been originally proposed to select the 198 independent variables that are relevant for a statistical analysis based on bivariate correlations 199 between them and one dependent variable. Herein they will be used to verify that all the considered 200 variables are significant for the performed analysis.

Finally, the calibrated computational map is evaluated by means of a multivariate computational susceptibility index, IS_{TCU}, which is assigned to each TCU according to the following formula:

203

$$204 IS_{TCU} = \sum_{i} W_{ik(i)} (4)$$

205

where: $W_{ik(i)}$ is the weight index of the relevant independent variable Vi related to the TCU belonging to class k(i) of that variable.

The deterministic analyses are based on TRIGRS (Montgomery and Dietrich, 1994), a physically-208 based model that is widely used for determining the distribution of shallow precipitation-induced 209 landslide source areas in different geological contexts (Godt et al., 2008; Montrasio et al., 2011; 210 Sorbino et al., 2010). The analysis couples a hydrologic model with an infinite slope stability 211 212 computation in order to analyze the pore water pressure regime and then evaluate a distribution of 213 factors of safety (FS) over large areas. In particular, TRIGRS models rainfall infiltration by solving the Richards equation for vertical infiltration that may occur during a precipitation event (Srivastava 214 215 and Yeh, 1991) in homogeneous isotropic materials. It analyzes the slope stability using a simple infinite-slope analysis (Taylor, 1948) to compute FS for each cell of the computational domain. 216 217 TRIGRS, used in conjunction with geographic information system (GIS) software, allows to obtain maps of FS results thus allowing to differentiate stable (FS>1) from unstable cells (FS≤1). Among 218 219 the disadvantages of TRIGRS, it is worth listing the following: steady or quasi-steady models are not applicable to several realistic cases (e.g. Wu and Sidle, 1995; Sorbino et al., 2010); the model 220 221 needs abundant and accurate spatial information over large portions of the study area in order to obtain reliable results; the results are quite sensitive to some of the input data, such as topographic 222 data (slope gradient), hydraulic properties of soils (saturated vertical hydraulic conductivity, 223 hydraulic diffusivity, saturated and residual volumetric water contents), initial water-table and soil 224

depths (Sorbino et al., 2010). Other input data the model needs are: rainfall data (with durations 225 ranging from hours to a few days), cohesion, friction angle and total unit weight of soils. To sum 226 up, realistic results from deterministic analyses based on TRIGRS are strictly linked to: a consistent 227 identification of the in situ pore pressure regime, a good knowledge of mechanical and hydraulic 228 soil properties, an accurate weathered thickness map, and a deep understanding of the triggering 229 mechanisms of the considered landslides. The sensitivity of TRIGRS to soil thickness has been 230 231 tested in Cascini et al. (2017). In the present paper we will test its sensitivity to the soil mechanical 232 properties and use it to properly identify the susceptibility computational map.

The performance of the susceptibility maps, obtained by both the statistical analysis and TRIGRS, is evaluated considering two indicators: the area under curve, AUC, of the related receiver operating characteristic curve (ROC) (Metz, 1978), and the overestimation index, OI, defined as:

236

$$0I = \frac{A_{out}}{A_{stab}} \cdot 100 \tag{5}$$

238
$$OI(5^{\circ}) = \frac{A_{out}}{A_{stab(5^{\circ})}} \cdot 100$$
(6)

239

where: A_{out} is the area computed unstable located outside the observed triggering area; A_{stab} is the area not affected by instability; $A_{stab(5^\circ)}$ is the area not affected by instability obtained excluding from the study area the zones with a slope gradient inferior or equal to 5°.

Concerning the AUC, a perfect model fitting would be characterised by an AUC value of 1 and a model not better than random would be characterised by an AUC value of 0.5. OI, originally defined "error index" by Sorbino et al. (2007, 2010), is proposed in order to individuate the "*overestimation*" area, intended here as the area computed as unstable by the model but not inventoried as area affected by landsliding.

248

249 3. Analyses and results

250 3.1. Statistical analyses

The variables employed within the statistical model have been expressed in raster format using 303071 square grid cells as TCUs, whose size is equal to 25 m². The dichotomous dependent variables, derived from two inventories of shallow landslides that occurred in 2009 and 2010, report 117 and 509 triggering areas which cover, respectively, 3421 and 10602 TCUs of the study area. The independent variables used in the analysis (Fig. 6, Table 1) are the five variables deemed to be

the most relevant following Calvello and Ciurleo (2016). The first three variables, i.e. elevation 256 zone (V1), slope gradient (V2) and slope curvature (V3), are computed from the native-resolution 257 variables, processing them by means of focal statistics techniques so that the information they carry 258 is averaged over a larger area around them. The area of influence considered for the three variables 259 depends on the diameter of the buffer considered around each TCU, herein called Dk, which is 260 respectively equal to 32 cells for V1, and to 4 cells for V2 and V3. All three variables are classified 261 according to a natural breaks criterion employing eight classes. The categorical variables 262 outcropping lithology (V4) is divided in four classes; in this case, as already mentioned, the 263 264 sandstone morphoselective scarp has been excluded from the analysis. The weathered rock thickness variable (V5) is divided in eight classes following the classification reported in the 265 266 employed thematic map.

Table 2, reporting the values of the statistical weights computed using Eq. (1) for each class of each 267 268 independent variable Vi, underlines that the maximum weights computed considering the dichotomous dependent variables, respectively dated 2009 and 2010, are both attributed to class 8 269 270 of variable V2 (W₂₈=1.28 in 2009 and W₂₈=1.15 in 2010). High values of weights are computed for variables V1, V3 and V5 both in 2009 and 2010: $W_{18}(2009)=0.98$; $W_{18}(2010)=0.68$; 271 272 $W_{31}(2009)=1.06$; $W_{31}(2010)=0.94$; $W_{52}(2009)=0.55$; $W_{52}(2010)=0.60$. Out of the four classes of variable V4, only one geological unit, which is the class corresponding to silty clays with sporadic 273 274 sand intercalations, assumes a positive weight value, $W_{44}=0.28$, for both years. Concerning low negative values of the weights, it is worth highlighting that only one class is characterised by a 275 weight value equal to -2.59 (W₂₃), while the W_{ij} equal to -3.00 is imposed whenever the argument 276 of the logarithm used in Eq. (1) is equal to zero (Ciurleo et al., 2016). 277

278 All the independent considered variables are relevant for the analysis as testified by the value of β_i 279 (Eq. 2) and σ_i (Eq. 3) respectively greater than 1.7 and 0.4 (Ciurleo et al., 2016), as reported in Table 3. The Table also shows the values of the two statistics needed to compute the bivariate 280 success index, TPR_i and FPR_i. Fig. 7 shows two landslide susceptibility computational maps 281 282 classified on the basis of the values assumed by the multivariate computational susceptibility index, 283 respectively considering the dichotomous dependent variables dated 2009 and 2010, IS_{TCU}, as follow: not susceptible, for IS_{TCU}≤0; susceptible for IS_{TCU}>0. The results indicate that about 20% of 284 285 the study area is susceptible (i.e. $IS_{TCU} > 0$) in 2009; this area increases to 24% in 2010. The success 286 of the analysis is verified by the high value assumed by the area under curve (AUC) of the ROC 287 curve plotted in the sensitivity versus (1 – specificity) space, with values ranging from 96% in 2009

to 95% in 2010. On the contrary, the values of OI and OI(5°) are never too low, as they range from OI=19% (in 2009) to OI=21% (in 2010) and OI(5°)=27% (in 2009) to OI(5°)=30% (in 2010).

290

291 3.2. Deterministic analyses

The input data employed within TRIGRS, as for the variables used in the statistical analyses, are expressed in raster format using 5×5 m² square grid cells as TCUs. The input data used in TRIGRS are the following: DEM, flow direction, thickness map, initial water table location, geotechnical properties of the soils and rainfall data.

Flow direction has been directly derived by DEM with single cell size equal to $5 \times 5 \text{ m}^2$. Following Cascini et al. (2017) it was possible to both reconstruct the thickness map (Fig. 5b) and localize the initial water table at the contact between intact and weathered rocks.

Referring to the geotechnical properties, the study area is analysed following two different cases 299 (Tab. 4): case 1, assigning the same geotechnical input data for the entire study area; case 2, 300 301 differentiating the geotechnical properties in relation to the thickness of the weathered rock. In particular, for case 1 the assigned geotechnical parameters are equal to the minimum values 302 available from the geotechnical input data. For case 2, different geotechnical properties are assumed 303 for three classes of thickness (0.5-1.0 m; 1.0-2.0 m and 2.0-3.0 m) corresponding to three different 304 305 shallow landslide mechanisms. These values have been provided by Cascini et al. (2017) who, by means of 2D geotechnical analyses, were able to back-analyse these shallow precipitation-induced 306 307 landslides and to provide different values of cohesion and friction angle for which the factor of 308 safety is equal to 1 along the assumed sliding surface. This is essentially due to the genesis of the 309 landslide mechanisms and to their close relationship with the weathering processes (Cascini et al., 310 2017).

The obtained results (Fig. 8) are provided for both cases assuming a value of rainfall input data 311 312 equal to 162 mm in six days (Cascini et al., 2017). The results clearly show a consistent decrease of 313 the unstable areas passing from case 1 (minimum geotechnical data) to case 2 (different 314 geotechnical data depending on the weathered thickness). This reduction appears evident also when comparing, using the indicators AUC, OI and OI(5°), the obtained results with the landslide 315 316 inventories dated 2009 and 2010. For Case 1, the AUC values range from 92% to 91%, respectively computed considering the shallow landslide inventories dated 2009 and 2010; the OI values range 317 from 18% (in 2009) to 16% (in 2010); the OI(5°) values range from 25% (in 2009) to 23 % (in 318 2010). For case 2, the obtained results show: AUC = 93%, OI = 9% and OI(5°)=13% for the 319

landslide inventory dated 2009; AUC= 92%, OI = 8% and OI(5°)=11% for the landslide inventory dated 2010. Comparing the results obtained from case 1 and case 2, it is evident that, despite the AUC value does not improve significantly, there is a significant reduction of the OI value, passing from 18% to 9% in 2009 and from 16% to 8% in 2010, and of the OI(5°) values, going from 25% to 13% (in 2009) and from 23% to 11% (in 2010).

325

326 4. Discussion and concluding remarks

The paper presented a comparison, within a study area located in southern Italy, between statistical 327 and deterministic methods used as tools for shallow landslides susceptibility assessment in clayey 328 soils. To this purpose, the information value method for the statistical analyses and TRIGRS for the 329 330 deterministic analyses have been used at large scale (1:5000). The information value method has been implemented using two dichotomous dependent variables (landslide inventories, years 2009 331 and 2010); TRIGRS model has been applied referring to two different geotechnical datasets (case 1 332 and case 2). In total, we obtained four landslide computational maps that have been compared by 333 means of three indicators: AUC of the ROC curves, OI and OI(5°). 334

335 The very high values attained by the AUC testify the success of all the performed analyses. Indeed, 336 independently from the method and the reference landslide inventory used, AUC is always greater than 90%, with values ranging from 95-96% for the statistical analyses and 91-92% for the 337 338 deterministic analyses. The 90% AUC value is defined by Swets (1988) and Fressard et al. (2014) as a threshold to overcome for reaching the best class of accuracy for a model being tested. In 339 340 particular Swets (1988) states that values between about 0.7 and 0.9 represent accuracies that are useful for some purposes, and higher values represent a rather high accuracy; whereas Fressard et 341 342 al. (2014) state that values between 0.7 and 0.8 reflect a fair performance of the model, values between 0.8 and 0.9 can be considered good, and values above 0.9 can be considered excellent. To 343 344 stress the relevance of the results obtained in this case study it is important to highlight that, within 345 the literature dealing with landslide susceptibility and hazard assessment, few analyses report AUC values higher than 85% (e.g., Akgun, 2012; Yesilnacara and Topalb, 2005; Lee and Pradhan, 2006; 346 Blahut et al., 2010; Yeon et al., 2010; Sterlacchini et al., 2011; Das et al., 2012) and even fewer 347 AUC values higher than 90% (e.g., Lee et al., 2008; Bui et al, 2012; Marjanović 2013; Nefeslioglu 348 349 et al., 2008; Ciurleo et al., 2016).

The success of the analyses is also related to the values assumed by over-prediction indexes such as: the modified success rate, MSR (Huang and Kao, 2006); the false positive rate of the contingency

table (Godt et al., 2008; Montrasio et al., 2011; Raia et al., 2013; Ciurleo et al., 2016); the error 352 index, EI (Sorbino et al., 2010); the landslide potential, LP (Vieira et al., 2010); and the false alarm 353 ratio, FAR (Liao et al., 2011). Among them, we choose the EI index, originally proposed by 354 Sorbino et al. (2007, 2010) and named OI in this paper, in order to quantify the overestimation 355 areas. In the case study, the area computed as unstable by the statistical analyses is equal to 1.5 Km² 356 (2009) and 1.8 Km² (2010), while unstable TRIGRS cells involve an area of 1.4 Km² (case 1) and 357 0.8 Km^2 (case 2). These values, if compared with the inventoried ones (0.1 Km² of triggering areas 358 inventoried in 2009 and 0.2 Km² in 2010), show a clear over-prediction ratio. This overestimation, 359 identified by means of OI and OI(5°), is more evident for the statistical analyses and progressively 360 decrease, moving from case 1 to case 2, when using TRIGRS. In particular, the statistical OI values 361 362 are equal to 19% in 2009 and 21% in 2010, while the TRIGRS OI values range from 16% to 18% for case 1 and from 8% to 9% for case 2. The difference in the OI values becomes more evident 363 364 when comparing the lowermost statistical OI(5°) value, equal to 27% (considering the 2009 landslides) with the lowermost $OI(5^{\circ})$ value for the analyses with TRIGRS, equal to 11% (for case 365 366 2 in 2010). To explain these findings, it is worthwhile highlighting herein that TRIGRS's case 1 refers to a simplified geotechnical scheme (low strength characteristics for the overall investigated 367 area) while case 2 employs different geotechnical parameters in relation to different thicknesses of 368 the weathered rock layer. In the latter case, very low values of both OI and $OI(5^{\circ})$ testify that an 369 exhaustive understanding of the main physical and mechanical characteristics of the analysed 370 phenomena is extremely relevant for the application of a deterministic model. Finally, considering 371 that the same common input data have been used (in terms of quality and accuracy) for the 372 statistical and the deterministic analyses, it is reasonable to assume that the higher overestimation 373 recorded by the statistical analyses is related to the inherent inability of statistical methods to 374 375 consider the physical processes within the soils predisposing or triggering the landslides.

In conclusion, the analyses performed over the same study area, to evaluate both the potentialities of the adopted methods and the consistency of the obtained results, indicate that statistical and deterministic analyses can be confidently developed at large scale for shallow landslide susceptibility assessment. The main results indicate that the statistical method provides shallow landslide susceptibility maps that are more conservative than those obtained by the physically-based model TRIGRS. The reliability of this latter progressively increases when the predisposing and triggering factors of the studied phenomena are properly characterized and used as input data.

383

REFERENCES

Antronico, L., Borrelli, L., Coscarelli, R., Pasqua, A.A., Petrucci, O., Gulla`, G., 2013. Slope movements induced by rainfalls damaging an urban area: The Catanzaro case study (Calabria, southern Italy). Landslides 10, 801–814.

Akgun, 2012. A comparison of landslide susceptibility maps produced by logistic regression, multicriteria decision, and likelihood ratio methods: a case study at İzmir, Turkey. Landslides 9, 93-106. Armas, I., Stroia, F., Giurgea, L., 2013. Statistic Versus deterministic Method for Landslide Susceptibility Mapping. Margottini et al. (eds.), Landslide Science and Practice, Vol. 3, DOI

10.1007/978-3-642-31310-3_52.

Baeza, C., Corominas, J., 2001. Assessment of shallow landslide susceptibility by means of multivariate statistical techniques. Earth Surface Processes and Landforms, 26, 1251–1263.

Baum, R.L., Coe, J.A., Godt, J.W., Harp, E.L., Reid, M.E., Savage, W.Z., Schulz, W.H., Brien D.L., Chleborad, A.F., McKenna, J. P., Michael J.A., 2005. Regional landslide-hazard assessment for Seattle, Washington, USA. Landslides 2, 266–279.

Blahut., J., VanWesten, C., Sterlacchini, S., 2010. Analysis of landslide inventories for accurate prediction of debris-flow source areas. Geophys J Roy Astron Soc, 119, 36–51.

Borrelli, L., Cofone, G., Coscarelli, R., Gullà, G., 2015. Shallow landslides triggered by consecutive rainfall events at Catanzaro strait (Calabria-Southern Italy). J. Maps 11, 730–744.

Borrelli, L., Gioffrè, D., Gullà, G., Moraci, N., 2012. Susceptibility to shallow and rapid landslides in ground alteration: A possible contribution of propagation modeling. Rendiconti Online Società Geologica Italiana, 21, 534-536.

Broothaerts, N., Kissi, E., Poesen, J., Van Rompaey, A., Getahun, K., Van Ranst, E., Diels, J., 2012. Spatial patterns, causes and consequences of landslides in the Gilgel Gibe catchment, SW Ethiopia. Catena 97, 127–136.

Bui, D., Pradhan, B., Lofman, O., Dick, O., 2012. Landslide susceptibility assessment in the Hoa Binh Province of Vietnam: a comparison of the Levenberg-Marquardt and Bayesian regularized neural networks. Geophys J Roy Astron Soc., 171–172, 12–29.

Calvello, M., Ciurleo, M., 2016. Optimal use of thematic maps for landslide susceptibility assessment by means of statistical analyses: case study of shallow landslides in fine grained soils. Proc. ISL 2016, Landslides and Engineered Slopes. Experience – Theory and Practice, Napoli, Italy, Vol. 2, p.537-544, ISBN: 978-1-138-02988-0.

Calvello, M., Cascini, L., Mastroianni, S., 2013. Landslide zoning over large areas from a sample inventory by means of scale-dependent terrain units. Geomorphology 182, 33–48.

This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 13)

Carrara, A., 1983. A multivariate model for landslide hazard evaluation. Mathematical Geology 15, 403–426.

Cascini, L., 2008. Applicability of landslide susceptibility and hazard zoning at different scales. Eng. Geol. 102, 164–177.

Cascini, L., Ciurleo, M., Di Nocera, S., Gullà, G., 2015. A new–old approach for shallow landslide analysis and susceptibility zoning in fine-grained weathered soils of southern Italy. Geomorphology 241, 371–381.

Cascini, L., Ciurleo, M., Di Nocera, S., 2017. Soil depth reconstruction for the assessment of the susceptibility to shallow landslides in fine-grained slopes. Landslides 14, 459–471.

Cervi, F., Berti, M., Borgatti, L., Ronchetti, F., Manenti, F., Corsini, A., 2010. Comparing predictive capability of statistical and deterministic methods for landslide susceptibility mapping: a case study in the northern Apennines (Reggio Emilia Province, Italy). Landslides 7, 433–444.

Ciurleo, M., Calvello, M., Cascini, L., 2016. Susceptibility zoning of shallow landslides in fine grained soils by statistical methods. Catena 139, 250–264.

Conforti, M., Robustelli, G., Muto, F., Critelli, S., 2012. Application and validation of bivariate GIS-based landslide susceptibility assessment for the Vitravo river catchment (Calabria, south Italy). Nat Hazards 61, 127–141.

Das, I., Stein, A., Kerle, N., Dadhwal, V., 2012. Landslide susceptibility mapping along road corridors in the Indian Himalayas using Bayesian logistic regression models. Geophys J Roy Astron Soc. 179, 116–125.

Fell, R., Corominas, J., Bonnard, C., Cascini, L., Leroi, E., Savage,W.Z., on behalf of the JTC-1 Joint Technical Committee on Landslides and Engineered Slopes, 2008. Guidelines for landslide susceptibility, hazard and risk zoning for land-use planning. Eng. Geol. 102, 85–98.

Frattini, P., Crosta, G., Carrara, A., 2010. Techniques for evaluating the performance of landslide susceptibility models. Engineering Geology 111, 62–72.

Fressard, M., Thiery, Y., Maquaire, O., 2014. Which data for quantitative landslide susceptibility mapping at operational scale: Case study of the Pays d'Auge plateau hillslopes (Normandy, France). Nat. Hazards Earth Syst. Sci. 14, 569–588.

Godt, J. W., Baum, R. B., Savage, W. Z., Salciarini, D., Schulz, W. H., and Harp, E. L., 2008. Transient deterministic shallow landslide modeling: Requirements for susceptibility and hazard assessments in a GIS framework, Eng. Geol. 102, 214–226.

This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 14)

Gullà, G., Aceto, L., Niceforo, D., 2004. Geotechnical characterisation of fine-grained soils affected by soil slips. Proc. of the 9th International Symposium on Landslides, Rio de Janeiro, Brazil, pp. 663–668.

Gullà, G., Aceto, L., Critelli, S., Perri, F., 2008. Geotechnical and mineralogical characterization of fine grained soils affected by soil slips. Proc. of the 10th International Symposium on Landslides, Xi'an, China, pp. 373–379.

Hadmoko, D.S., Lavigne, F., Samodra, G., 2017. Application of a semiquantitative and GIS-based statistical model to landslide susceptibility zonation in Kayangan Catchment, Java, Indonesia. Nat Hazards. DOI 10.1007/s11069-017-2772-z.

Huang, J. C., Kao, S. J., 2006. Optimal estimator for assessing landslide model performance Hydrol. Earth Syst. Sci. 10, 957–965.

Hungr, O., Evans, S.G., Bovis, M.J., Hutchinson, J.N., 2001. A review of the classification of landslides of the flow type. Environmental and Engineering Geoscience 3, 221-238.

O. Hungr, S. Leroueil, L. Picarelli (2014). The Varnes classification of landslide types, an update. Landslides, 11: 167–194

Iovine, G.G.R., Greco, R., Gariano, S.L., Pellegrino, A.D., Terranova, O.G. 2014. Shallowlandslide susceptibility in the Costa Viola mountain ridge (southern Calabria, Italy) with considerations on the role of causal factors. Nat Hazards 73, 111–136.

Jenks, G. F., 1977. Optimal data classification for choropleth maps. Geography Department Occasional Paper No. 2 2, University of Kansas, Lawrence, KS, 1977.

Kavzoglu, T., Sahin, E.K., Colkesen, I., 2015. Selecting optimal conditioning factors in shallow translational landslide susceptibility mapping using genetic algorithm. Engineering Geology 192, 101–112.

Lee, J.H, and Park, H.J., 2016. Assessment of shallow landslide susceptibility using the transient infiltration flow model and GIS-based probabilistic approach. Landslides 13, 885–903.

Lee, S., Pradhan, B., 2006. Probabilistic landslide hazard and risk mapping on Penang Island, Malaysia. J Earth Syst Sci. 115, 661–672.

Lee, C., Huang, C., Lee, J., Pan, K., Lin, M., Dong, J., 2008. Statistical approach to storm eventinduced landsliZ des susceptibility. Nat Hazards Earth Syst Sci. 8, 941–960.

Leventhal, A.R., Kotze, G.P., 2008. Landslide susceptibility and hazard mapping in Australia for land-use planning — with reference to challenges in metropolitan suburbia. Engineering Geology 102, 238–250.

Liao, Z., Hong, Y., Kirschbaum, D., Adler, R. F., Gourley, J. J., and Wooten, R., 2011. Evaluation of TRIGRS (transient rainfall infiltration and grid-based regional slope-stability analysis)'s predictive skill for hurricane-triggered landslides: a case study in Macon County, North Carolina, Nat. Hazards. 58, 325–339.

Marjanović, 2013. Comparing the performance of different landslide susceptibility models in ROC space. Landslide Science and Practice, Volume 1: Landslide inventory and susceptibility and hazard zoning, 579-584.

Martinović, K., Gavina, K., Reale, C., 2016. Development of a landslide susceptibility assessment for a rail network. Engineering Geology 215, 1–9.

McDougall, S., Hungr, O., 2004. A model for the analysis of rapid landslide run out motion across three dimensional terrain. Canadian Geotechnical Journal 41, 1084–1097.

Meisina, C., 2006. Characterisation of weathered clayey soils responsible for shallow landslides. Nat. Hazards Earth Syst. Sci. 6, 825–838.

Metz, C.E., 1978. Basic principles of ROC analysis. Semin. Nucl. Med. 8, 283–298.

Montgomery, D.R., Dietrich, W.E., 1994. A physically based model for the topographic control on shallow landsliding. Water Resour. Res. 30, 1153–1171.

Montrasio, L., Valentino, R., and Losi, G. L., 2011. Towards a real-time susceptibility assessment of rainfall-induced shallow landslides on a regional scale, Nat. Hazards Earth Syst. Sci. 11, 1927–1947, doi:10.5194/nhess-11-1927-2011.

Nandi, A., Shakoor, A., 2009. A GIS-based landslide susceptibility evaluation using bivariate and multivariate statistical analyses. Engineering Geology 110, 11–20.

Nefeslioglu, HA., Gokceoglu, C., Sonmezb, H., 2008. An assessment on the use of logistic regression and artificial neural networks with different sampling strategies for the preparation of landslide susceptibility maps. Eng. Geol. 97, 171–191.

Park, H.J., Lee, J.H., Woo, I, 2013. Assessment of rainfall-induced shallow landslide susceptibility using a GIS-based probabilistic approach. Engineering Geology 161, 1–15.

Raia, S., Alvioli, M., Rossi, M., Baum, R. L., Godt, J. W., and Guzzetti, F., 2013. Improving predictive power of physically based rainfall-induced shallow landslide models: a probabilistic approach, Geosci. Model Dev. Discuss., 6, 1367–1426, doi:10.5194/gmdd-6-1367-2013.

Regmi, N.R., Giardino, J.R., McDonald, E.V., Vitek, J.D., 2014. A comparison of logistic regression-based models of susceptibility to landslides in western Colorado, USA. Landslides 11, 247–262.

Reza, M., Daneshvar, M., 2014. Landslide susceptibility zonation using analytical hierarchy process and GIS for the Bojnurd region, northeast of Iran. Landslides 11, 1079–1091.

Romer, C., Ferentinou, M., 2016. Shallow landslide susceptibility assessment in a semiarid environment — A Quaternary catchment of KwaZulu-Natal, South Africa. Engineering Geology 201, 29–44.

Salciarini, D, Godt, J.W., Savage, W.Z., 2006. Modeling regional initiation of rainfall-induced shallow landslides in the eastern Umbria Region of central Italy. Landslides 3, 181–194.

Soeters, R., van Westen, C.J., 1996. Slope instability recognition, analysis and zonation. In: Turner, A.K., Schuster, R.L. (Eds.), Landslides Investigation and Mitigation. TRB Special Report 247. National Academy Press, Washington D.C., pp. 129–177.

Sorbino, G., Sica, C., Cascini, L., Cuomo, S., 2007. On the forecasting of flowslides triggering areas using physically based models. Proceedings of 1st North American Landslides Conference, AEG Special Publication. 23, pp. 305–315.

Sorbino, G., Sica, C., Cascini, L., 2010. Susceptibility analysis of shallow landslides source areas using physically based models. Nat. Hazards 53, 313–332.

Srivastava, R., Yeh, T.-C. J., 1991. Analytical solutions for one-dimensional, transient infiltration toward the water table in homogeneous and layered soils: Water Resources Research v. 27, pp. 753–762.

Sterlacchini, S., Ballabio, C., Blahut, J., Masetti, M., Sorichetta, A., 2011. Spatial agreement of predicted patterns in landslide susceptibility maps. Geophys J Roy Astron Soc. 125, 51–61

Swets, J.A., 1988. Measuring the accuracy of diagnostic systems. Science 240, 1285–1293.

Tangestani, M.H., 2009. A comparative study of Dempster-Shafer and fuzzy models for landslide susceptibility mapping using a GIS: an experience from Zagros Mountains, SW Iran. J. Asian Earth Sci. 35, 66–73.

Taylor, DW, 1948. Fundamentals of soil mechanics. Wiley, New York, 700 p.

Vieira, B.C., Fernandes, N.F., Filho, O.A., 2010. Shallow landslide prediction in the Serra do Mar, São Paulo, Brazil, Nat. Hazards Earth Syst. Sci., 10, 1829–1837, doi:10.5194/nhess-10-1829-2010.

Wu, W., Sidle, RC., 1995. A distributed slope stability model for steep forested basins. Water Resour Res 31, 2097–2110.

Yeon, Y., Han, J., Ryu, K., 2010. Landslide susceptibility mapping in Injae, Korea using a decision tree. Eng Geol. 166, 274–28.

Yerro, A., Alonso, E.E., Pinyol, N.M., 2016. Run-out of landslides in brittle soils. Computers and Geotechnics 80, 427–439.

This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 17)

Yesilnacara, E., Topalb, T., 2005. Landslide susceptibility mapping: A comparison of logistic regression and neural networks methods in a medium scale study, Hendek region (Turkey). Eng Geol. 79, 251–266.

Yin, K.J., Yan, T.Z., 1988. Statistical prediction model for slope instability of metamorphosed rock. Proc. 5th Int.l Symposium on Landslides, Lausanne, Switzerland vol. 2l, pp. 1269–1272.

| Class | V1 elevation zone | V2 slope gradient (°) | V3 slope curvature | V4 geological unit | V5 weathered |
|-------|-------------------|-----------------------|--------------------|--|---------------|
| | (m) | | (m-1) | (-) | thickness (-) |
| 1 | 30.39 to 55.86 | 0.00 to 4.12 | - 6.12 to - 2.06 | Colluvial deposits | 0.00-0.50 |
| 2 | 55.87 to 81.33 | 4.13 to 8.59 | - 2.05 to - 1.02 | Halluvial deposits | 0.50-1.00 |
| 3 | 81.34 to 105.98 | 8.60 to 12.89 | - 1.01 to - 0.36 | Slope debris | 1.00-1.50 |
| 4 | 105.99 to 130.63 | 12.90 to 17.00 | - 0.35 to 0.13 | Silty clays with sporadic sand intercalation | 1.50–2.00 |
| 5 | 130.64 to 154.46 | 17.01 to 20.94 | 0.14 to 0.74 | | 2.00-3.00 |
| 6 | 154.47 to 178.29 | 20.95 to 25.06 | 0.75 to 1.56 | | 3.00-4.00 |
| 7 | 178.30 to 230.76 | 25.07 to 30.07 | 1.57 to 2.71 | | 4.00-5.00 |
| 8 | 230.77 to 239.91 | 30.07 to 45.64 | 2.72 to 7.87 | | > 5.00 |

Tab.1. Classification of the independent variables employed in the statistical analysis at large scale.

Tab.2. Weights assigned to the independent variables in the statistical analysis at large scale.

| Weights | V1 | V2 | V3 | V4 | V5 | Weights | V1 | V2 | V3 | V4 | V5 |
|---------|--------|--------|--------|--------|--------|---------|--------|--------|--------|--------|--------|
| 2009 | | | | | | 2010 | | | | | |
| Wi1 | - 3.00 | - 3.00 | 1.06 | - 3.00 | - 0.26 | Wi1 | - 3.00 | - 3.00 | 0.94 | - 3.00 | - 0.26 |
| Wi2 | - 3.00 | - 3.00 | 0.55 | - 1.24 | 0.55 | Wi2 | - 3.00 | - 3.00 | 0.50 | - 1.38 | 0.60 |
| Wi3 | - 3.00 | - 3.00 | 0.04 | - 0.40 | 0.27 | Wi3 | - 3.00 | - 2.59 | 0.05 | - 0.43 | 0.17 |
| Wi4 | - 3.00 | - 1.52 | - 0.55 | 0.28 | - 0.10 | Wi4 | - 1.29 | - 1.26 | - 0.52 | 0.28 | - 0.17 |
| Wi5 | - 0.46 | - 0.52 | - 0.11 | | - 0.80 | Wi5 | 0.07 | - 0.22 | - 0.13 | | - 0.52 |
| Wi6 | 0.25 | 0.23 | 0.13 | | - 3.00 | Wi6 | 0.42 | 0.36 | 0.22 | | - 1.42 |
| Wi7 | 0.63 | 0.92 | 0.23 | | - 3.00 | Wi7 | 0.66 | 0.89 | 0.40 | | - 3.00 |
| Wi8 | 0.98 | 1.28 | 0.25 | | - 3.00 | Wi8 | 0.68 | 1.15 | 0.48 | | - 3.00 |

| Tab.3. Values of parameters and inc | xes needed to select th | he independent variables relevant f | for the |
|--------------------------------------|-------------------------|-------------------------------------|---------|
| statistical analysis at large scale. | | | |

| Variables | 2009 | | | | 2010 | | | | |
|-----------|------------------|------------------|------|--------------|------------------|------------------|------|------|----------|
| | TPR _i | FPR _i | β | σ_{i} | TPR _i | FPR _i | β | σ | Relevant |
| V1 | 95.8% | 22.0% | 4.36 | 2.20 | 99.3% | 32.5% | 3.05 | 2.17 | Yes |
| V2 | 95.7% | 17.6% | 5.45 | 2.45 | 91.3% | 15.8% | 5.78 | 2.45 | Yes |
| V3 | 73.3% | 34.6% | 2.12 | 0.77 | 72.9% | 33.7% | 2.17 | 0.74 | Yes |
| V4 | 93.2% | 48.8% | 1.91 | 0.85 | 93.9% | 47.7% | 1.97 | 0.86 | Yes |
| V5 | 84.0% | 31.0% | 2.71 | 2.23 | 84.0% | 29.7% | 2.83 | 2.22 | Yes |

This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 19)

Tab.4. The geotechnical input data used for TRIGRS analyses.

| | | TRIGRS | – Case 1 | | | | | | |
|----------------------|--------------------|------------------------------|--------------|----------------|----------------------------|--|--|--|--|
| Unit weight y | Effective cohesion | Friction angle \(\phi'\) (°) | Soil depth h | Hydraulic | Diffusivity D | | | | |
| (kN/m ³) | c' (kPa) | | TRIGRS (m) | conductivity K | TRIGRS (m ² /s) | | | | |
| | | | | (m/s) | | | | | |
| 18 | 3 | 27 | variable | 5.00E - 06 | 2.87-05 | | | | |
| TRIGRS – Case 2 | | | | | | | | | |
| Unit weight y | Effective cohesion | Friction angle φ' (°) | Soil depth h | Hydraulic | Diffusivity D | | | | |
| (kN/m ³) | c' (kPa) | | TRIGRS (m) | conductivity K | TRIGRS (m ² /s) | | | | |
| | | | | (m/s) | | | | | |
| 18 | 3 | 27 | 0.5-1.0 | 5.00E - 06 | 2.87E - 05 | | | | |
| 18 | 5 | 27 | 1.0-2.0 | 5.00E - 07 | 3.49E - 05 | | | | |
| 18 | 8 | 27 | 2.0-3.0 | 5.00E - 07 | 3.49E - 05 | | | | |

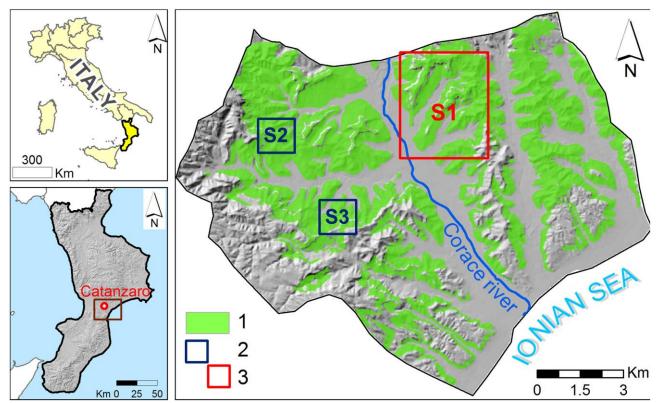
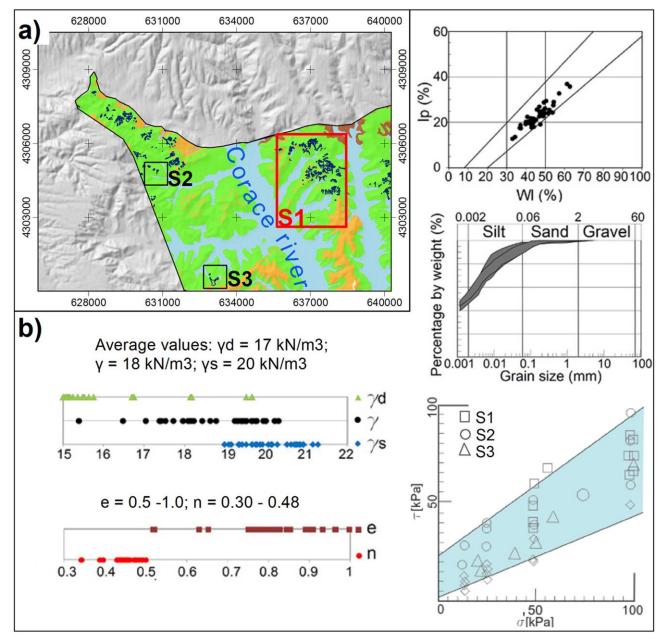


Fig.1. Geographical location of the study area. Legend: 1. Pliocene light blue-grey silty clays; 2. zones frequently involved by shallow landslides studied in the past (S2 and S3); 3. study area (S1).

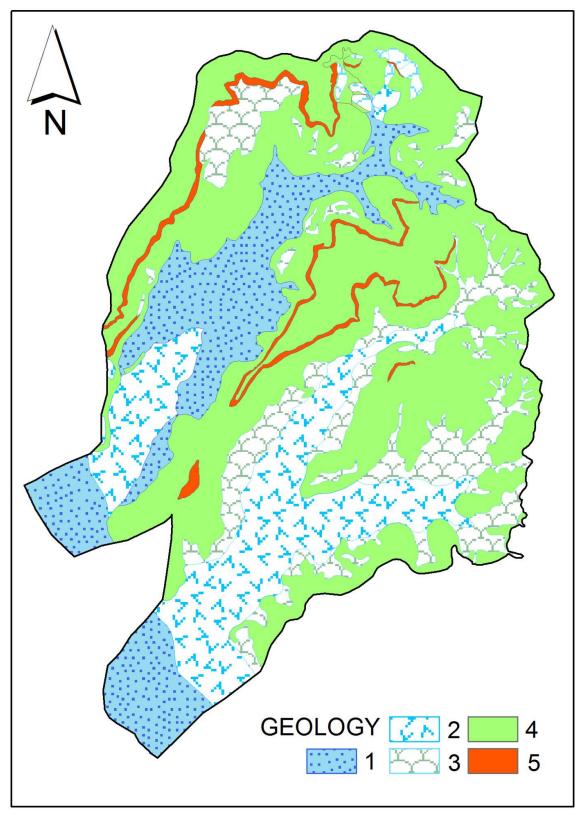
This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 21)

Fig. 2. Available in situ investigations and laboratory tests (modified from Cascini et al., 2017). a) Spatial location. b) Physical and mechanical properties of weathered clays. From the top to the bottom: plasticity chart, grain size distribution envelope, index properties (dry unit weight - γd (kN/m3), natural unit weight - γ (kN/m3), the saturated unit weight - γs (kN/m3), index void - e (-), soil porosity - n (-)), shear strength (modified from Gullà et al. 2008; Cascini et al. 2015).



This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 22)

Fig. 3. Lithological map of the study area (S1), Legend: 1. alluvial deposits; 2. colluvial; 3. landslide debris; 4. light blue-grey silty clays; 5. Sandstone.



This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 23)

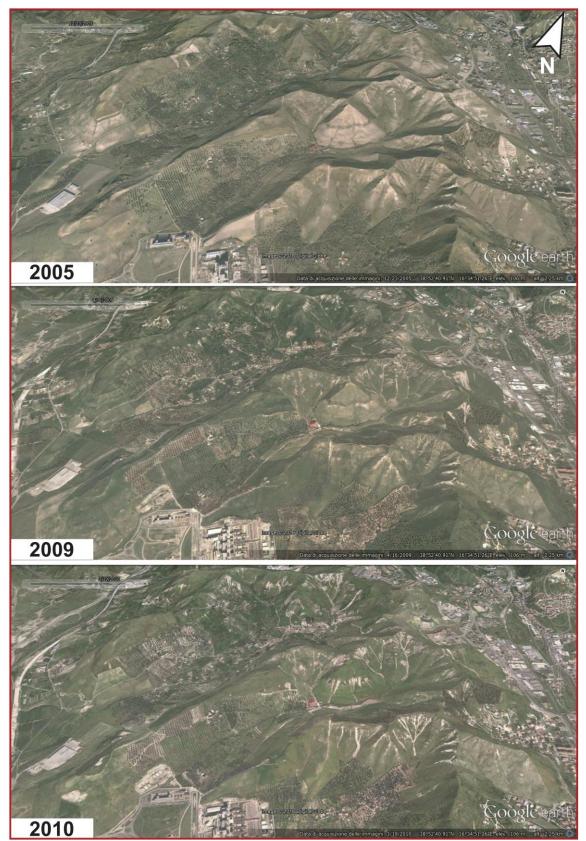


Fig. 4. 3D view from Google Earth of December 2005, April 2009 and March 2010 for the study area (S1).

This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 24)

Fig. 5. Thematic map employed for the analysis at S1: a) Landslide inventories. Legend: landslide source areas from 2009 (1) and 2010 (2) inventories; b) Weathered rock thickness map.

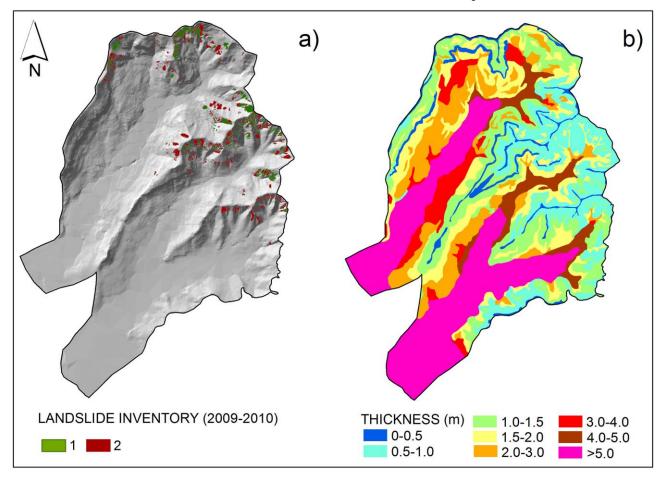
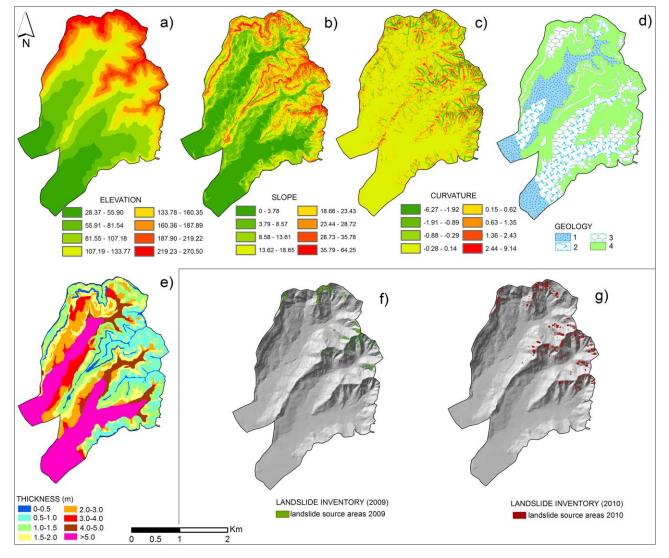


Fig. 6. Variables employed in the statistical analysis. Independent variables: (a) elevation zones, V1; (b) slope gradient, V2; (c) slope curvature, V3; (d) geology V4;(e) weathered rock thickness, V5. Dependent variables: (f) landslide source areas from 2009; g) landslide source areas from 2010. Legend for (d): 1. alluvial deposits; 2. colluvial; 3. slope debris; 4. light blue-grey silty clays.



This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 26)

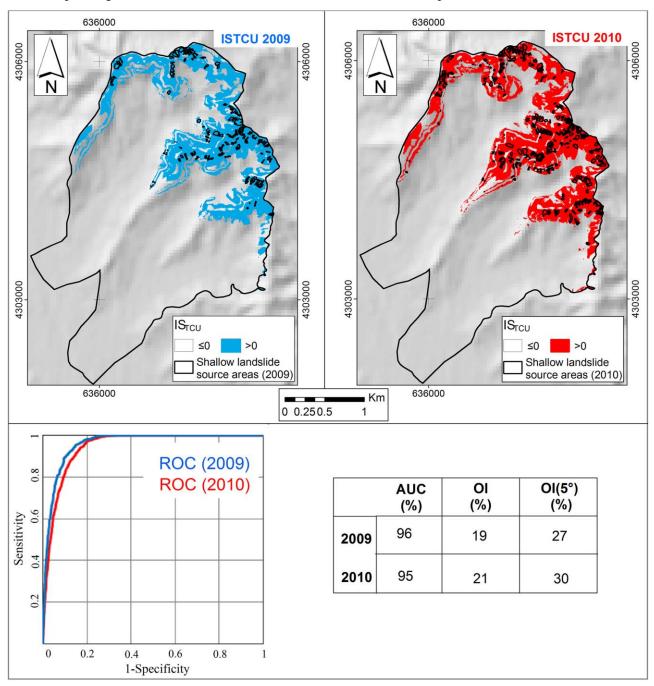
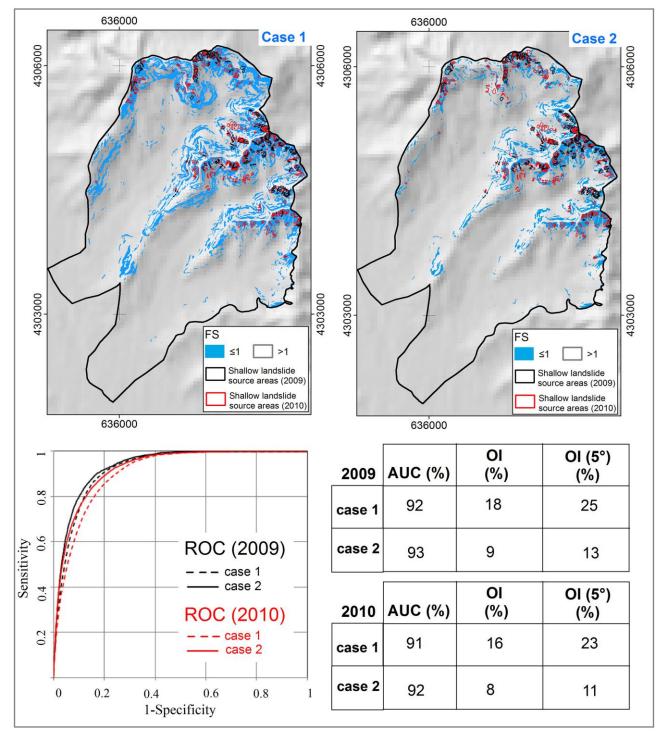


Fig. 7. Results of the statistical analyses at 2009 and 2010: landslide susceptibility computational maps; receiver operating characteristic curves; values of statistical indicators of performance AUC, OI and OI(5°).

This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 27)

Fig. 8. Results of the TRIGRS analyses at 2009 and 2010 for case 1 and case 2: landslide susceptibility computational maps; receiver operating characteristic curves; values of statistical indicators of performance AUC, OI and $OI(5^{\circ})$.



This is a post-peer-review, pre-copyedit version of an article published in "Engineering geology". The final authenticated version is available at: https://www.sciencedirect.com/science/article/pii/S0013795216308419 (page 28)